The Spectrum of Strong Behavioral Equivalences for Nondeterministic and Probabilistic Processes

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We present a spectrum of trace-based, testing, and bisimulation equivalences for nondeterministic and probabilistic processes whose activities are all observable. For every equivalence under study, we examine the discriminating power of three variants stemming from three approaches that differ for the way probabilities of events are compared when nondeterministic choices are resolved via deterministic schedulers. We show that the first approach – which compares two resolutions relatively to the probability distributions of all considered events – results in a fragment of the spectrum compatible with the spectrum of behavioral equivalences for fully probabilistic processes. In contrast, the second approach – which compares the probabilities of the events of a resolution with the probabilities of the same events in possibly different resolutions – gives rise to another fragment composed of coarser equivalences that exhibits several analogies with the spectrum of behavioral equivalences for fully nondeterministic processes. Finally, the third approach – which only compares the extremal probabilities of each event stemming from the different resolutions – yields even coarser equivalences that, however, give rise to a hierarchy similar to that stemming from the second approach.

1 Introduction

Process algebras are mathematically rigorous languages that have been widely used to model and analyze the behavior of interacting systems. Their structural operational semantics associates with each process term a labeled transition system (LTS), whose states are the terms themselves and whose labels are the actions that each term can perform. In order to abstract from unwanted details, the operational semantics is often coupled with observational mechanisms that permit equating those systems that cannot be distinguished by external entities. The resulting behavioral equivalences heavily depend on how the specified systems are expected to be used. Indeed, there is still disagreement on which are the "reasonable" observations and how their outcomes can be used to distinguish or identify systems. Thus, many equivalences have been proposed and much work has been done to assess their discriminating power and mutual relationships.

The first study in this direction was done by [9]. There, most of the then known equivalences over LTS models were "ordered" and it was shown that *trace equivalences* (equating systems performing the same sequences of actions) are strictly coarser than *decorated-trace equivalences* (equating systems performing the same sequences of actions and refusing/accepting the same sets of actions after them), which in turn are strictly coarser than *bisimulation equivalences* (equating systems performing the same sequences of actions and recursively exhibiting the same behavior). It was also shown that the equivalence obtained by *testing* processes with external observers was coincident with *failure equivalence*

obtained via traces decorated with refusal sets. Afterwards, [14] built the first spectrum that relates twelve different equivalences and set up a general testing scenario that could be used to generate many more equivalences.

When process algebras have been enriched with additional dimensions to deal with probabilistic, stochastic, and timed systems, new behavioral equivalences have been defined and possible classifications have been proposed. Here, we would like to concentrate on equivalences for probabilistic systems. For this class of systems, comparative results have been obtained only for so-called fully probabilistic systems [20, 17, 2] or only for bisimulation and testing relations [2, 22, 29, 34].

In this paper, we aim at a systematic account of the known probabilistic equivalences for nondeterministic *and* probabilistic systems and introduce, motivate, and relate some new ones. We shall consider an extension of the LTS model combining nondeterminism and probability that we call NPLTS, in which every action-labeled transition goes from a source state to a probability distribution over target states rather than to a single target state [21, 25]. Actions will be assumed to be visible (i.e., we shall not admit τ -actions) and, for the considered strong equivalences, resolutions of nondeterminism will be derived by applying memoryless deterministic (as opposed to randomized) schedulers.

When defining behavioral relations over NPLTS models, the idea is to compare resolutions on the basis of the probabilities of *equivalence-specific events*, like (i) performing certain sequences of actions, (ii) exhibiting certain decorated traces, or (iii) reaching certain sets of equivalent states via given actions.

The typical approach followed in the literature (see, e.g., [28, 26, 27]) consists of comparing the *probability distributions of all equivalence-specific events* of two resolutions. Two processes are considered as equivalent if, for each resolution of any of the two processes, there exists a resolution of the other process such that the probability of *each* equivalence-specific event is the same in the two resolutions (*fully matching resolutions*). For the known relations based on this approach, we have that the probabilistic bisimilarity in [28] implies the probabilistic failure equivalence in [27] that in turn implies the probabilistic trace equivalence in [26]. All these relations are conservative extensions of the corresponding relations defined over fully nondeterministic models [16, 6] and fully probabilistic models [13, 20, 17], but in many situations they turn out to have a high discriminating power.

A different approach has been followed in the literature for defining testing equivalences (see, e.g., [35, 19, 27, 11]). Instead of comparing individual resolutions of the parallel composition of processes and tests, the comparison is performed between the *extremal probabilities* of reaching success *over all resolutions* generated by the experiments on processes under test (*max-min-matching resolution sets*). In this case, it holds that the resulting probabilistic testing equivalence is implied by the probabilistic bisimilarity in [28], but it is related neither to the probabilistic failure equivalence in [27] nor to the probabilistic trace equivalence in [26] when restricting attention to deterministic schedulers. Moreover, the resulting probabilistic testing equivalence subsumes testing equivalence for fully probabilistic processes [7], but it is not a conservative extension of testing equivalence for fully nondeterministic processes [10].

Recently, in [8, 32, 31, 3, 5] a further approach has appeared that compares resolutions on the basis of the *probabilities of individual equivalence-specific events*. Thus, a resolution of any of the two processes can be matched, with respect to *different equivalence-specific events*, by *different resolutions* of the other process (*partially matching resolutions*). For the behavioral relations resulting from this approach, which weakens the impact of schedulers, we have that probabilistic bisimilarity implies probabilistic failure equivalence, which in turn implies probabilistic testing equivalence, which finally implies probabilistic trace equivalence. This approach has contributed to the development of new probabilistic bisimilarities in [8], [32, 5], and [31] that, unlike the one in [28], are characterized by standard probabilistic logics such as quantitative μ -calculus, PML, and PCTL/PCTL*, respectively. Moreover, in the case of testing equivalence this approach has the advantage of being conservative also for fully nonde-

terministic models [3], while in the case of trace equivalence it surprisingly results in a congruence with respect to parallel composition (full version of [3]).

In our view, the motivations behind the three approaches outlined above are all very reasonable. Indeed, when applied to fully nondeterministic processes or fully probabilistic processes, they give rise to well-studied relations that for the fully nondeterministic setting fit into the spectra in [9, 14] and for the fully probabilistic setting fit into the spectra in [20, 17]. The situation is significantly different when the three approaches are instantiated for nondeterministic *and* probabilistic processes, as in that case they give rise to a much wider variety of relations.

In this paper, we study the relationships between the equivalences for NPLTS models that stem from the three approaches. For each approach, we consider the three main families of equivalences, namely trace-based, testing, and bisimulation equivalences. To the best of our knowledge, this is the first comparative study of different kinds of behavioral equivalences over models featuring both nondeterministic and probabilistic aspects. Such a study is even more on demand after the recent introduction of new equivalences, like the ones in [8, 32, 31, 3, 5], that have interesting properties.

To have a full picture of the spectrum, the reader is referred to Fig. 4 in the concluding section. There, the equivalences stemming from the same approach are contained in boxes with the same shape (hexagonal, rounded, or rectangular) and the equivalences specifically introduced for the purposes of this paper are in dashed boxes. We would like to stress that the original contribution of the paper is not given by the equivalences that we introduce to fill in gaps, *but is the spectrum itself*.

We shall see that the family of equivalences that assign a central role to schedulers by requiring fully matching resolutions, yields a hierarchy that is in accordance with the one for fully probabilistic processes in [20, 17]. Conversely, the family of equivalences that assign a weaker role to schedulers by requiring partially matching resolutions, gives rise to relations that are coarser than the former and yields a hierarchy that is in accordance with the one for fully nondeterministic processes in [9, 14]. Finally, the family of equivalences that only consider extremal probabilities, has again several analogies with the fully nondeterministic spectrum and yields even coarser equivalences. There are however some noticeable anomalies in the last two families, given by a few equivalences suffering from isolation.

The rest of the paper is organized as follows. In Sect. 2, we introduce the NPLTS model. In Sects. 3 to 5, we define and compare, respectively, the trace-based, testing, and bisimulation equivalences that arise from the three approaches outlined above. Finally, in Sect. 6 we draw some conclusions and graphically summarize the results by depicting the spectrum of all the considered equivalences.

2 Nondeterministic and Probabilistic Processes

Processes combining nondeterminism and probability are typically described by means of extensions of the LTS model, in which every action-labeled transition goes from a source state to a *probability distribution over target states* rather than to a single target state. They are essentially Markov decision processes and are representative of a number of slightly different probabilistic computational models including internal nondeterminism such as, e.g., concurrent Markov chains [33], alternating probabilistic models [15, 35, 24], probabilistic automata in the sense of [25], and the denotational probabilistic models in [18] (see [30] for an overview). We formalize them as a variant of simple probabilistic automata [25].

Definition 2.1 A nondeterministic and probabilistic labeled transition system, NPLTS for short, is a triple (S,A,\longrightarrow) where S is an at most countable set of states, A is a countable set of transition-labeling actions, and $\longrightarrow \subseteq S \times A \times Distr(S)$ is a transition relation with Distr(S) being the set of discrete probability distributions over S.

A transition (s, a, \mathcal{D}) is written $s \xrightarrow{a} \mathcal{D}$. We say that $s' \in S$ is not reachable from s via that a-transition if $\mathcal{D}(s') = 0$, otherwise we say that it is reachable with probability $p = \mathcal{D}(s')$. The reachable states form the support of $\mathcal{D}: supp(\mathcal{D}) = \{s' \in S \mid \mathcal{D}(s') > 0\}$. We write $s \xrightarrow{a}$ to indicate that s has an a-transition. The choice among all the transitions departing from s is nondeterministic, while the choice of the target state for a specific transition is probabilistic. An NPLTS represents (i) a *fully nondeterministic process* when every transition leads to a distribution that concentrates all the probability mass into a single state or (ii) a *fully probabilistic process* when every state has at most one outgoing transition.

An NPLTS can be depicted as a directed graph-like structure in which vertices represent states and action-labeled edges represent action-labeled transitions. Given a transition $s \stackrel{a}{\longrightarrow} \mathscr{D}$, the corresponding a-labeled edge goes from the vertex representing state s to a set of vertices linked by a dashed line, each of which represents a state $s' \in supp(\mathscr{D})$ and is labeled with $\mathscr{D}(s')$ – label omitted if $\mathscr{D}(s')$ = 1. Figure 1 shows eighteen NPLTS models, nine of which are fully nondeterministic.

In this setting, a computation is a sequence of state-to-state steps, each denoted by $s \xrightarrow{a} s'$ and derived from a state-to-distribution transition. Formally, given an NPLTS $\mathscr{L} = (S,A,\longrightarrow)$ and $s,s' \in S$, we say that $c \equiv s_0 \xrightarrow{a_1} s_1 \xrightarrow{a_2} s_2 \dots s_{n-1} \xrightarrow{a_n} s_n$ is a computation of \mathscr{L} of length n from $s = s_0$ to $s' = s_n$ iff for all $i = 1, \dots, n$ there exists a transition $s_{i-1} \xrightarrow{a_i} \mathscr{D}_i$ such that $s_i \in supp(\mathscr{D}_i)$, with $\mathscr{D}_i(s_i)$ being the execution probability of step $s_{i-1} \xrightarrow{a_i} s_i$ conditioned on the selection of transition $s_{i-1} \xrightarrow{a_i} \mathscr{D}_i$ of \mathscr{L} at state s_{i-1} . We denote by first(c) and first(c) the initial state and the final state of c, respectively, and by $\mathscr{C}_{fin}(s)$ the set of finite-length computations from s.

We call resolution of s any possible way of resolving nondeterminism starting from s. Each resolution is a tree-like structure whose branching points represent probabilistic choices. This is obtained by unfolding from s the graph structure underlying \mathcal{L} and by selecting at each state a single transition of \mathcal{L} (deterministic scheduler) or a convex combination of equally labeled transitions of \mathcal{L} (randomized scheduler) among all the transitions possible from that state. Below, we introduce the notion of resolution arising from a deterministic scheduler as a fully probabilistic NPLTS. Notice that, when \mathcal{L} is fully nondeterministic, resolutions boil down to computations.

Definition 2.2 Let $\mathcal{L} = (S, A, \longrightarrow)$ be an NPLTS and $s \in S$. We say that an NPLTS $\mathcal{L} = (Z, A, \longrightarrow_{\mathcal{L}})$ is a resolution of s obtained via a deterministic scheduler iff there exists a state correspondence function $corr_{\mathcal{L}}: Z \to S$ such that $s = corr_{\mathcal{L}}(z_s)$, for some $z_s \in Z$, and for all $z \in Z$ it holds that:

- If $z \xrightarrow{a}_{\mathscr{Z}} \mathscr{D}$, then $corr_{\mathscr{Z}}(z) \xrightarrow{a} \mathscr{D}'$ with $\mathscr{D}(z') = \mathscr{D}'(corr_{\mathscr{Z}}(z'))$ for all $z' \in Z$.
- If $z \xrightarrow{a_1}_{\mathscr{Z}} \mathscr{D}_1$ and $z \xrightarrow{a_2}_{\mathscr{Z}} \mathscr{D}_2$, then $a_1 = a_2$ and $\mathscr{D}_1 = \mathscr{D}_2$.

We denote by Res(s) the set of resolutions of s and by $Res_{\max}(s)$ the set of maximal resolutions of s, i.e., the resolutions of s that cannot be further extended in accordance with the graph structure of \mathscr{L} and the constraints above. Since $\mathscr{L} \in Res(s)$ is fully probabilistic, the probability prob(c) of executing $c \in \mathscr{C}_{\text{fin}}(z_s)$ can be defined as the product of the (no longer conditional) execution probabilities of the individual steps of c, with prob(c) being always equal to 1 if \mathscr{L} is fully nondeterministic. This notion is lifted to $C \subseteq \mathscr{C}_{\text{fin}}(z_s)$ by letting $prob(C) = \sum_{c \in C} prob(c)$ whenever none of the computations in C is a proper prefix of one of the others.

We finally introduce a notion of fully synchronous parallel composition for NPLTS models that is instrumental to the definition of testing equivalences.

Definition 2.3 Let $\mathcal{L}_i = (S_i, A, \longrightarrow_i)$ be an NPLTS for i = 1, 2. The parallel composition of \mathcal{L}_1 and \mathcal{L}_2 is the NPLTS $\mathcal{L}_1 \parallel \mathcal{L}_2 = (S_1 \times S_2, A, \longrightarrow)$ where $\longrightarrow \subseteq (S_1 \times S_2) \times A \times Distr(S_1 \times S_2)$ is such that $(s_1, s_2) \stackrel{a}{\longrightarrow} \mathcal{D}$ iff $s_1 \stackrel{a}{\longrightarrow} 1 \mathcal{D}_1$ and $s_2 \stackrel{a}{\longrightarrow} 2 \mathcal{D}_2$ with $\mathcal{D}(s'_1, s'_2) = \mathcal{D}_1(s'_1) \cdot \mathcal{D}_2(s'_2)$ for each $(s'_1, s'_2) \in S_1 \times S_2$.

3 Trace-Based Equivalences for NPLTS Models

Trace-based equivalences examine the probability with which two states perform computations labeled with the same (decorated) traces for each possible way of resolving nondeterminism. As outlined in Sect. 1, there are three different approaches to defining them. The first approach is to match resolutions according to *trace-based distributions*, which means that for each resolution of one of the two states there must exist a resolution of the other state such that, *for every (decorated) trace*, the two resolutions have the same probability of performing a computation labeled with that (decorated) trace. In other words, matching resolutions of the two states are related by the fully probabilistic version of the trace-based equivalence (fully matching resolutions). The second approach is to consider *a single (decorated) trace at a time*, i.e., to anticipate the quantification over (decorated) traces with respect to the quantification over resolutions. In this way, differently labeled computations of a resolution of one of the two states are allowed to be matched by computations of several different resolutions of the other state (partially matching resolutions). The third approach is to compare only the *extremal probabilities* of performing each (decorated) trace over the various resolutions (max-min-matching resolution sets).

We say that a computation is compatible with a trace $\alpha \in A^*$ iff the sequence of actions labeling its steps is equal to α . Given an NPLTS $\mathscr{L} = (S,A,\longrightarrow)$, $s \in S$, and $\mathscr{L} \in Res(s)$, we denote by $\mathscr{CC}(z_s,\alpha)$ the set of α -compatible computations in $\mathscr{C}_{fin}(z_s)$ and by $Res_{\alpha}(s)$ the set of resolutions in Res(s) having no computations corresponding to proper prefixes of α -compatible computations of \mathscr{L} . In each of the following definitions, we assume $s_1, s_2 \in S$ and we explicitly add a reference whenever the defined equivalence has already appeared in the literature. In some definitions, we indicate with \Box/\Box the supremum/infimum of a set of numbers in $\mathbb{R}_{[0,1]}$ and we assume it to be 0 when the set is empty.

Definition 3.1 (*Probabilistic trace-distribution equivalence* $- \sim_{PTr,dis} - [26]$)

 $s_1 \sim_{\mathsf{PTr.dis}} s_2$ iff for each $\mathscr{Z}_1 \in Res(s_1)$ there exists $\mathscr{Z}_2 \in Res(s_2)$ such that for all $\alpha \in A^*$:

$$prob(\mathscr{CC}(z_{s_1},\alpha)) = prob(\mathscr{CC}(z_{s_2},\alpha))$$

and symmetrically for each $\mathscr{Z}_2 \in Res(s_2)$.

Definition 3.2 (*Probabilistic trace equivalence* $-\sim_{PTr}$ – [3])

 $s_1 \sim_{\mathsf{PTr}} s_2$ iff for all $\alpha \in A^*$ it holds that for each $\mathscr{Z}_1 \in Res(s_1)$ there exists $\mathscr{Z}_2 \in Res(s_2)$ such that:

$$prob(\mathscr{CC}(z_{s_1},\alpha)) = prob(\mathscr{CC}(z_{s_2},\alpha))$$

and symmetrically for each $\mathscr{Z}_2 \in Res(s_2)$.

Definition 3.3 (*Probabilistic* $\Box \Box$ -trace equivalence $-\sim_{PTr. \Box \Box}$)

 $s_1 \sim_{\mathsf{PTr}, \sqcup \sqcap} s_2$ iff for all $\alpha \in A^*$:

$$\bigsqcup_{\mathcal{Z}_1 \in Res_{\alpha}(s_1)} prob(\mathscr{CC}(z_{s_1}, \alpha)) = \bigsqcup_{\mathcal{Z}_2 \in Res_{\alpha}(s_2)} prob(\mathscr{CC}(z_{s_2}, \alpha))$$

$$\bigsqcup_{\mathcal{Z}_1 \in Res_{\alpha}(s_1)} prob(\mathscr{CC}(z_{s_1}, \alpha)) = \bigsqcup_{\mathcal{Z}_2 \in Res_{\alpha}(s_2)} prob(\mathscr{CC}(z_{s_2}, \alpha))$$

$$\mathscr{Z}_1 \in Res_{\alpha}(s_1)$$

A variant that additionally considers completed computations was introduced in the literature of fully nondeterministic models in order to equip trace equivalence with deadlock sensitivity. We denote by $\mathscr{CC}(z_s,\alpha)$ the set of completed α -compatible computations from z_s . Each of these computations c belongs to $\mathscr{CC}(z_s,\alpha)$ and is such that $corr_{\mathscr{Z}}(last(c))$ has no outgoing transitions in \mathscr{L} .

Definition 3.4 (*Probabilistic completed-trace-distribution equivalence* $-\sim_{PCTr,dis}$)

 $s_1 \sim_{\text{PCTr,dis}} s_2$ iff for each $\mathscr{Z}_1 \in Res(s_1)$ there exist $\mathscr{Z}_2, \mathscr{Z}'_2 \in Res(s_2)$ such that for all $\alpha \in A^*$:

$$prob(\mathscr{CC}(z_{s_1},\alpha)) = prob(\mathscr{CC}(z_{s_2},\alpha))$$

 $prob(\mathscr{CC}(z_{s_1},\alpha)) = prob(\mathscr{CCC}(z'_{s_2},\alpha))$

and symmetrically for each $\mathscr{Z}_2 \in Res(s_2)$.

Definition 3.5 (*Probabilistic completed-trace equivalence* $-\sim_{PCTr}$)

 $s_1 \sim_{\text{PCTr}} s_2 \text{ iff } \underline{\text{for all } \alpha \in A^*} \text{ it holds that for each } \mathscr{Z}_1 \in Res(s_1) \text{ there exist } \mathscr{Z}_2, \mathscr{Z}_2' \in Res(s_2) \text{ such that:}$

$$prob(\mathscr{CC}(z_{s_1},\alpha)) = prob(\mathscr{CC}(z_{s_2},\alpha))$$
$$prob(\mathscr{CC}(z_{s_1},\alpha)) = prob(\mathscr{CCC}(z'_{s_2},\alpha))$$

and symmetrically for each $\mathscr{Z}_2 \in Res(s_2)$.

Definition 3.6 (*Probabilistic* $\sqcup \sqcap$ -completed-trace equivalence $- \sim_{PCTr, \sqcup \sqcap}$)

 $s_1 \sim_{\text{PCTr}, \square} s_2 \text{ iff for all } \alpha \in A^*$:

$$\begin{array}{ccc} \bigsqcup\limits_{\mathscr{Z}_{1} \in Res_{\alpha}(s_{1})} prob(\mathscr{CC}(z_{s_{1}},\alpha)) &=& \bigsqcup\limits_{\mathscr{Z}_{2} \in Res_{\alpha}(s_{2})} prob(\mathscr{CC}(z_{s_{2}},\alpha)) \\ \bigsqcup\limits_{\mathscr{Z}_{1} \in Res_{\alpha}(s_{1})} prob(\mathscr{CC}(z_{s_{1}},\alpha)) &=& \prod\limits_{\mathscr{Z}_{2} \in Res_{\alpha}(s_{2})} prob(\mathscr{CC}(z_{s_{2}},\alpha)) \end{array}$$

and:

$$\bigsqcup_{\mathcal{Z}_1 \in Res_{\alpha}(s_1)} prob(\mathscr{CCC}(z_{s_1}, \alpha)) = \bigsqcup_{\mathcal{Z}_2 \in Res_{\alpha}(s_2)} prob(\mathscr{CCC}(z_{s_2}, \alpha))$$

$$\prod_{\mathcal{Z}_1 \in Res_{\alpha}(s_1)} prob(\mathscr{CCC}(z_{s_1}, \alpha)) = \prod_{\mathcal{Z}_2 \in Res_{\alpha}(s_2)} prob(\mathscr{CCC}(z_{s_2}, \alpha))$$

Definition 3.7 (*Probabilistic failure-distribution equivalence* $-\sim_{PF,dis}$ – [27])

 $s_1 \sim_{\text{PF,dis}} s_2$ iff for each $\mathscr{Z}_1 \in Res(s_1)$ there exists $\mathscr{Z}_2 \in Res(s_2)$ such that for all $\varphi \in A^* \times 2^A$:

$$prob(\mathscr{FCC}(z_{s_1},\varphi)) = prob(\mathscr{FCC}(z_{s_2},\varphi))$$

and symmetrically for each $\mathscr{Z}_2 \in Res(s_2)$.

Definition 3.8 (*Probabilistic failure equivalence* $-\sim_{PF}$ – [3])

 $s_1 \sim_{PF} s_2$ iff for all $\varphi \in A^* \times 2^A$ it holds that for each $\mathscr{Z}_1 \in Res(s_1)$ there exists $\mathscr{Z}_2 \in Res(s_2)$ such that:

$$\overline{prob}(\mathscr{FCC}(z_{s_1},\varphi)) = prob(\mathscr{FCC}(z_{s_2},\varphi))$$

and symmetrically for each $\mathscr{Z}_2 \in Res(s_2)$.

Definition 3.9 (*Probabilistic* $\sqcup \sqcap$ -failure equivalence $- \sim_{PF, \sqcup \sqcap}$)

 $s_1 \sim_{\mathrm{PF}, \sqcup \sqcap} s_2$ iff for all $\varphi = (\alpha, F) \in A^* \times 2^A$:

$$\Box prob(\mathscr{FCC}(z_{s_1},\varphi)) = \Box prob(\mathscr{FCC}(z_{s_2},\varphi))
\Box prob(\mathscr{FCC}(z_{s_1},\varphi)) = \Box prob(\mathscr{FCC}(z_{s_2},\varphi))
\Box prob(\mathscr{FCC}(z_{s_1},\varphi)) = \Box prob(\mathscr{FCC}(z_{s_2},\varphi))
\mathscr{Z}_1 \in Res_{\alpha}(s_1)$$

Definition 3.10 (*Probabilistic failure-trace-distribution equivalence* $-\sim_{PFTr.dis}$)

Same as Def. 3.7 with $\phi \in (A \times 2^A)^*$ and \mathscr{FSCC} in place of $\varphi \in A^* \times 2^A$ and \mathscr{FCC} , respectively.

Definition 3.11 (*Probabilistic failure-trace equivalence* $-\sim_{PFTr}$)

Same as Def. 3.8 with $\phi \in (A \times 2^A)^*$ and \mathscr{FSCC} in place of $\varphi \in A^* \times 2^A$ and \mathscr{FCC} , respectively.

Definition 3.12 (*Probabilistic* $\sqcup \sqcap$ -failure-trace equivalence $- \sim_{PFTr, \sqcup \sqcap}$)

Same as Def. 3.9 with $\phi \in (A \times 2^A)^*$ and \mathscr{FSCC} in place of $\varphi \in A^* \times 2^A$ and \mathscr{FCC} , respectively.

A different generalization towards liveness properties is readiness semantics. A ready pair is an element $\varrho \in A^* \times 2^A$ formed by a trace ϱ and a decoration R called ready set. We say that ϱ is compatible with ϱ iff $\varrho \in \mathscr{CC}(z_s, \varrho)$ and the set of actions labeling the transitions in \mathscr{L} departing from $corr_{\mathscr{L}}(last(c))$ is precisely ϱ . We denote by $\mathscr{RCC}(z_s, \varrho)$ the set of ϱ -compatible computations from z_s . Moreover, we call ready trace an element $\varrho \in (A \times 2^A)^*$ given by a sequence of $\varrho \in \mathbb{N}$ pairs of the form (a_i, R_i) . We say that $\varrho \in \mathscr{C}_{fin}(z_s)$ is compatible with $\varrho \in \mathscr{CC}(z_s, a_1 \dots a_n)$ and, denoting by $\varrho \in \mathscr{L}_{fin}(z_s)$ the state reached by $\varrho \in \mathscr{L}_{fin}(z_s)$ is precisely $\varrho \in \mathscr{L}_{fin}(z_s)$ is precisely $\varrho \in \mathscr{L}_{fin}(z_s)$ the set of actions labeling the transitions in \mathscr{L} departing from $corr_{\mathscr{L}}(z_s)$ is precisely $\varrho \in \mathscr{L}_{fin}(z_s)$. We denote by $\mathscr{RCC}(z_s, \varrho)$ the set of ϱ -compatible computations from $\varrho \in \mathscr{L}_{fin}(z_s)$ is precisely $\varrho \in \mathscr{L}_{fin}(z_s)$.

Definition 3.13 (*Probabilistic readiness-distribution equivalence* $-\sim_{PR.dis}$)

$$s_1 \sim_{\mathsf{PR},\mathsf{dis}} s_2$$
 iff for each $\mathscr{Z}_1 \in Res(s_1)$ there exists $\mathscr{Z}_2 \in Res(s_2)$ such that for all $\varrho \in A^* \times 2^A$: $prob(\mathscr{RCC}(z_{s_1},\varrho)) = prob(\mathscr{RCC}(z_{s_2},\varrho))$

and symmetrically for each $\mathscr{Z}_2 \in Res(s_2)$.

Definition 3.14 (*Probabilistic readiness equivalence* $-\sim_{PR}$)

 $s_1 \sim_{PR} s_2$ iff for all $\varrho \in A^* \times 2^A$ it holds that for each $\mathscr{Z}_1 \in Res(s_1)$ there exists $\mathscr{Z}_2 \in Res(s_2)$ such that:

$$\overline{prob}(\mathscr{RCC}(z_{s_1},\varrho)) = prob(\mathscr{RCC}(z_{s_2},\varrho))$$

and symmetrically for each $\mathscr{Z}_2 \in Res(s_2)$.

Definition 3.15 (*Probabilistic* $\sqcup \sqcap$ -readiness equivalence $- \sim_{PR, \sqcup \sqcap}$)

 $s_1 \sim_{\text{PR}, \sqcup \sqcap} s_2$ iff for all $\varrho = (\alpha, R) \in A^* \times 2^A$:

$$\mathcal{Z}_{1} \in Res_{lpha}(s_{1}) \qquad \mathcal{Z}_{2} \in Res_{lpha}(s_{2}) \qquad \qquad \mathcal{Z}_{2} \in Res_{lpha}(s_{2}) \qquad \qquad \mathcal{Z}_{2} \in Res_{lpha}(s_{2}) \qquad \qquad \mathcal{Z}_{3} \in Res_{lpha}(s_{2}) \qquad \qquad \mathcal{Z}_{4} \in Res_{lpha}(s_{2}) \qquad \qquad \mathcal{Z}_{5} \in Res_{lpha}(s_{2}) \qquad \qquad \mathcal{Z}_{$$

Definition 3.16 (*Probabilistic ready-trace-distribution equivalence* $-\sim_{PRTr,dis}$)

Same as Def. 3.13 with $\rho \in (A \times 2^A)^*$ and \mathscr{RTCC} in place of $\varrho \in A^* \times 2^A$ and \mathscr{RCC} , respectively.

Definition 3.17 (*Probabilistic ready-trace equivalence* $-\sim_{PRTr}$)

Same as Def. 3.14 with $\rho \in (A \times 2^A)^*$ and \mathscr{RTCC} in place of $\varrho \in A^* \times 2^A$ and \mathscr{RCC} , respectively.

Definition 3.18 (*Probabilistic* $\Box \Box$ -ready-trace equivalence $-\sim_{PRTr, \Box \Box}$)

Same as Def. 3.15 with
$$\rho \in (A \times 2^A)^*$$
 and \mathscr{RTCC} in place of $\rho \in A^* \times 2^A$ and \mathscr{RCC} , respectively.

The eighteen trace-based equivalences defined above are all backward compatible with the corresponding trace-based equivalences respectively defined in [6, 23] for fully nondeterministic processes and in [20, 17] for fully probabilistic processes.

Theorem 3.19 Let $\sigma \in \{\text{RTr}, \text{FTr}, R, F, \text{CTr}, \text{Tr}\}$ with $\sim_{P\sigma, \text{dis}}$, $\sim_{P\sigma}$, and $\sim_{P\sigma, \sqcup \sqcap}$ being the equivalences defined above, $\sim_{\sigma, \text{fnd}}$ being the corresponding equivalence defined for fully nondeterministic processes, and $\sim_{\sigma, \text{fpr}}$ being the corresponding equivalence defined for fully probabilistic processes. Then:

- 1. $\sim_{P\sigma,dis} = \sim_{P\sigma} = \sim_{P\sigma, \sqcup \sqcap} = \sim_{\sigma,find}$ over fully nondeterministic NPLTS models.
- 2. $\sim_{P\sigma,dis} = \sim_{P\sigma} = \sim_{P\sigma, \sqcup \sqcap} = \sim_{\sigma,fpr}$ over fully probabilistic NPLTS models.

We now investigate the relationships among the eighteen trace-based equivalences. As expected, each equivalence relying on trace-based distributions is finer than the corresponding equivalence considering a single (decorated) trace at a time, which in turn is finer than the corresponding equivalence based on extremal probabilities of (decorated) traces. For the equivalences of the first type, similar to the fully probabilistic spectrum in [20, 17] it turns out that the readiness semantics coincides with the failure semantics. In contrast, for the other two types of equivalences, unlike the fully nondeterministic spectrum in [14] no connection can be established between readiness semantics and failure semantics.

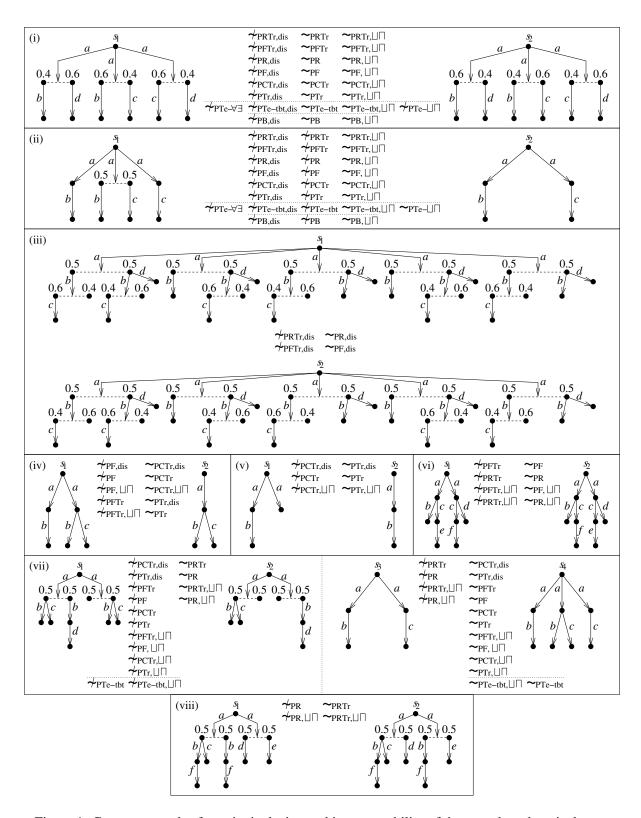


Figure 1: Counterexamples for strict inclusion and incomparability of the trace-based equivalences

Theorem 3.20 It holds that:

- 1. $\sim_{\pi.\mathrm{dis}} \subseteq \sim_{\pi} \subseteq \sim_{\pi.\sqcup\sqcap}$ for all $\pi \in \{\mathsf{PRTr}, \mathsf{PFTr}, \mathsf{PR}, \mathsf{PF}, \mathsf{PCTr}, \mathsf{PTr}\}.$
- 2. $\sim_{PRTr,dis} = \sim_{PFTr,dis}$ over finitely-branching NPLTS models.
- 3. $\sim_{PR,dis} = \sim_{PF,dis}$ over finitely-branching NPLTS models.
- 4. $\sim_{PFTr dis} \subseteq \sim_{PF dis} \subseteq \sim_{PCTr dis} \subseteq \sim_{PTr dis}$
- 5. $\sim_{PFTr} \subseteq \sim_{PF} \subseteq \sim_{PCTr} \subseteq \sim_{PTr}$.
- 6. $\sim_{PFTr, |||} \subseteq \sim_{PF, |||} \subseteq \sim_{PCTr, ||||} \subseteq \sim_{PTr, ||||}$

All the inclusions above are strict, as shown in Figs. 1(i) to (vi). It is worth noting the isolation of \sim_{PRTr} , \sim_{PR} , $\sim_{PRTr, \sqcup \sqcap}$, and $\sim_{PR, \sqcup \sqcap}$, each of which is incomparable with $\sim_{PCTr, dis}$, $\sim_{PTr, dis}$, \sim_{PFTr} , \sim_{PF} , \sim_{PCTr} , \sim_{PTr} , $\sim_{PFTr, \sqcup \sqcap}$, $\sim_{PCTr, \sqcup \sqcap}$, and $\sim_{PTr, \sqcup \sqcap}$, as shown in Fig. 1(vii). Moreover, Figs. 1(i) and (iv) show that $\sim_{PFTr, \sqcup \sqcap}$, and $\sim_{PF, \sqcup \sqcap}$ are incomparable with $\sim_{PCTr, dis}$ and $\sim_{PTr, dis}$, while Figs. 1(ii) and (iv) show that $\sim_{PFTr, \sqcup \sqcap}$ and $\sim_{PFTr, \sqcup \sqcap}$ are also incomparable with \sim_{PCTr} and \sim_{PTr} . Finally, Figs. 1(vi) and (viii) show that $\sim_{PRTr, \sqcup \sqcap}$ and $\sim_{PRTr, \sqcup \sqcap}$ are incomparable with \sim_{PR} and $\sim_{PCTr, \sqcup \sqcap}$ is incomparable with $\sim_{PCTr, \sqcup \sqcap}$ is incomparable with $\sim_{PTr, dis}$, and Figs. 1(ii) and (v) show that $\sim_{PCTr, \sqcup \sqcap}$ is incomparable with \sim_{PTr} .

4 Testing Equivalences for NPLTS Models

Testing equivalences consider the probability of two processes of performing computations along which the same tests are passed. Tests specify which actions of a process are permitted at each step and, in this setting, can be formalized as NPLTS models equipped with a success state. For the sake of simplicity, we restrict ourselves to finite tests, each of which has finitely many states, finitely many outgoing transitions from each state, an acyclic graph structure, and hence finitely many computations leading to success.

Definition 4.1 A nondeterministic and probabilistic test, NPT for short, is a finite NPLTS $\mathscr{T} = (O, A, \longrightarrow)$ where O contains a distinguished success state denoted by ω that has no outgoing transitions. We say that a computation of \mathscr{T} is successful iff its last state is ω .

Definition 4.2 Let $\mathscr{L} = (S, A, \longrightarrow)$ be an NPLTS and $\mathscr{T} = (O, A, \longrightarrow_{\mathscr{T}})$ be an NPT. The interaction system of \mathscr{L} and \mathscr{T} is the NPLTS $\mathscr{I}(\mathscr{L}, \mathscr{T}) = \mathscr{L} \parallel \mathscr{T}$ where:

- Every element $(s, o) \in S \times O$ is called a configuration and is said to be successful iff $o = \omega$.
- A computation of $\mathscr{I}(\mathscr{L},\mathscr{T})$ is said to be successful iff its last configuration is successful. Given $s \in S$, $o \in O$, and $\mathscr{Z} \in Res(s,o)$, we denote by $\mathscr{SC}(z_{s,o})$ the set of successful computations from the state $z_{s,o}$ of \mathscr{Z} corresponding to the configuration (s,o) of $\mathscr{I}(\mathscr{L},\mathscr{T})$.

Due to the possible presence of equally labeled transitions departing from the same state, there is not necessarily a single probability value with which an NPLTS passes a test. Thus, given two states s_1 and s_2 of the NPLTS under test and the initial state o of the test, we need to compute the probability of performing a successful computation from the two configurations (s_1, o) and (s_2, o) in every maximal resolution of the interaction system. One option is comparing, for the two configurations, *only the extremal values of these success probabilities* over all maximal resolutions of the interaction system. An alternative option is comparing *all the success probabilities* and requiring that for each maximal resolution of either configuration there is a matching maximal resolution of the other configuration.

Definition 4.3 (*Probabilistic* $\sqcup \sqcap$ -testing equivalence $-\sim_{\text{PTe-}\sqcup\sqcap}$ – [35, 19, 27, 11])

 $s_1 \sim_{\mathsf{PTe-} \sqcup \sqcap} s_2$ iff for every NPT $\mathscr{T} = (O, A, \longrightarrow_{\mathscr{T}})$ with initial state $o \in O$:

$$\bigsqcup_{\mathcal{Z}_1 \in Res_{\max}(s_1,o)} prob(\mathscr{SC}(z_{s_1,o})) = \bigsqcup_{\mathcal{Z}_2 \in Res_{\max}(s_2,o)} prob(\mathscr{SC}(z_{s_2,o}))$$

$$\bigsqcup_{\mathcal{Z}_1 \in Res_{\max}(s_1,o)} prob(\mathscr{SC}(z_{s_1,o})) = \bigsqcup_{\mathcal{Z}_2 \in Res_{\max}(s_2,o)} prob(\mathscr{SC}(z_{s_2,o}))$$

$$\mathcal{Z}_1 \in Res_{\max}(s_1,o)$$

Definition 4.4 (*Probabilistic* $\forall \exists$ -testing equivalence $- \sim_{\text{PTe-} \forall \exists} - [3]$)

 $s_1 \sim_{\mathsf{PTe} \cdot \forall \exists} s_2$ iff for every NPT $\mathscr{T} = (O, A, \longrightarrow_{\mathscr{T}})$ with initial state $o \in O$ it holds that for each $\mathscr{Z}_1 \in Res_{\max}(s_1, o)$ there exists $\mathscr{Z}_2 \in Res_{\max}(s_2, o)$ such that:

$$prob(\mathscr{SC}(z_{s_1,o})) = prob(\mathscr{SC}(z_{s_2,o}))$$

and symmetrically for each $\mathcal{Z}_2 \in Res_{max}(s_2, o)$.

Neither $\sim_{\text{PTe-}\sqcup\sqcap}$ nor $\sim_{\text{PTe-}\forall\exists}$ is backward compatible with the testing equivalence defined in [10] for fully nondeterministic processes. For instance, Fig. 2(i) shows two such processes related by classical testing equivalence that are distinguished by $\sim_{\text{PTe-}\sqcup\sqcap}$ and $\sim_{\text{PTe-}\forall\exists}$. The reason of the higher discriminating power of the latter two equivalences arises from the presence of probabilistic choices within tests, which results in the capability of making copies of the process under test [1] and hence in an unrealistic estimation of success probabilities [12]. In order to counterbalance this strong discriminating power, as illustrated in [3] the idea is to consider *success probabilities in a trace-by-trace fashion* rather than on entire resolutions. Since traces come again into play, the idea can be implemented in three different ways by following the three approaches used in Sect. 3.

In the following, given a state s of an NPLTS, a state o of an NPT, and a trace $\alpha \in A^*$, we denote by $Res_{\max,\mathscr{C},\alpha}(s,o)$ the set of resolutions $\mathscr{Z} \in Res_{\max}(s,o)$ such that $\mathscr{CCC}(z_{s,o},\alpha) \neq \emptyset$, i.e., the maximal resolutions of $z_{s,o}$ having at least one completed α -compatible computation. Moreover, for each such resolution \mathscr{Z} , we denote by $\mathscr{SCC}(z_{s,o},\alpha)$ the set of successful α -compatible computations from $z_{s,o}$.

Definition 4.5 (Probabilistic trace-by-trace-distribution testing equivalence $-\sim_{\text{PTe-tbt.dis}}$)

 $s_1 \sim_{\text{PTe-tbt,dis}} s_2$ iff for every NPT $\mathscr{T} = (O, A, \longrightarrow_{\mathscr{T}})$ with initial state $o \in O$ it holds that for each $\mathscr{L}_1 \in Res_{\max}(s_1, o)$ there exists $\mathscr{L}_2 \in Res_{\max}(s_2, o)$ such that $\underline{\text{for all } \alpha \in A^*}$ it holds that $\mathscr{CCC}(z_{s_1, o}, \alpha) \neq \emptyset$ implies $\mathscr{CCC}(z_{s_2, o}, \alpha) \neq \emptyset$ and:

$$prob(\mathscr{SCC}(z_{s_1,o},\alpha)) = prob(\mathscr{SCC}(z_{s_2,o},\alpha))$$

and symmetrically for each $\mathscr{Z}_2 \in Res_{\max}(s_2, o)$.

Definition 4.6 (*Probabilistic trace-by-trace testing equivalence* $- \sim_{\text{PTe-tbt}} - [3]$)

 $s_1 \sim_{\mathsf{PTe-tbt}} s_2$ iff for every NPT $\mathscr{T} = (O, A, \longrightarrow_{\mathscr{T}})$ with initial state $o \in O$ and $\underline{\mathsf{for}} \ \mathsf{all} \ \alpha \in A^*$ it holds that for each $\mathscr{Z}_1 \in Res_{\max,\mathscr{C},\alpha}(s_1,o)$ there exists $\mathscr{Z}_2 \in Res_{\max,\mathscr{C},\alpha}(s_2,o)$ such that:

$$prob(\mathscr{SCC}(z_{s_1,o},\alpha)) = prob(\mathscr{SCC}(z_{s_2,o},\alpha))$$

and symmetrically for each $\mathscr{Z}_2 \in Res_{\max,\mathscr{C},\alpha}(s_2,o)$.

Definition 4.7 (*Probabilistic* $\Box \Box$ -trace-by-trace testing equivalence $-\sim_{\text{PTe-tbt}, \Box \Box}$)

 $s_1 \sim_{\mathsf{PTe-tbt}, \sqcup \sqcap} s_2$ iff for every NPT $\mathscr{T} = (O, A, \longrightarrow_{\mathscr{T}})$ with initial state $o \in O$ and for all $\alpha \in A^*$ it holds that $Res_{\max,\mathscr{C},\alpha}(s_1,o) \neq \emptyset$ iff $Res_{\max,\mathscr{C},\alpha}(s_2,o) \neq \emptyset$ and:

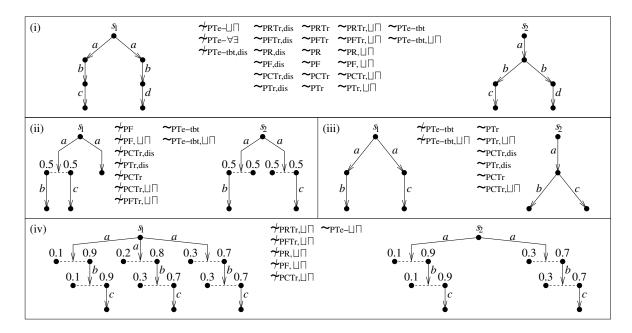


Figure 2: Counterexamples for strict inclusion and incomparability of the testing equivalences

While only $\sim_{\text{PTe-tbt}}$ and $\sim_{\text{PTe-tbt}, \sqcup \sqcap}$ are backward compatible with the testing equivalence defined in [10] for fully nondeterministic processes – which we denote by $\sim_{\text{Te,fnd}}$ (see Fig. 2(i) for the counterexamples) – all the five testing equivalences defined above are backward compatible with the testing equivalence defined in [7] for fully probabilistic processes – which we denote by $\sim_{\text{Te,fpr}}$.

Theorem 4.8 It holds that:

- 1. $\sim_{PTe-tbt} = \sim_{PTe-tbt, \sqcup \sqcap} = \sim_{Te,fnd}$ over fully nondeterministic NPLTS models.
- 2. $\sim_{\text{PTe-} \sqcup \sqcap} = \sim_{\text{PTe-} \forall \exists} = \sim_{\text{PTe-tbt}, \text{dis}} = \sim_{\text{PTe-tbt}} = \sim_{\text{PTe-tbt}, \sqcup \sqcap} = \sim_{\text{Te}, \text{fpr}} \text{ over fully probabilistic NPLTS models.}$

We now investigate the relationships of the five testing equivalences among themselves (first two properties below) and with the eighteen trace-based equivalences (last three properties below). It turns out that $\sim_{PTe-\forall\exists}$ and $\sim_{PTe-tbt,dis}$ perform exactly the same identifications. Unlike the fully nondeterministic spectrum – where the testing semantics coincides with the failure semantics when all actions are observable [9] – here $\sim_{PTe-tbt,dis}$ is finer than $\sim_{PFTr,dis}$ while $\sim_{PTe-tbt}$ and $\sim_{PTe-tbt, \Box\Box}$ are coarser than $\sim_{PTe-tdt}$ and $\sim_{PTe-tdt}$, respectively. In contrast, $\sim_{PTe-t\Box\Box}$ has no inclusion relationship with the failure semantics.

Theorem 4.9 It holds that:

- 1. $\sim_{\text{PTe-}\forall\exists} \subseteq \sim_{\text{PTe-}\sqcup\sqcap} \subseteq \sim_{\text{PTe-tbt},\sqcup\sqcap}$.
- 2. $\sim_{\text{PTe-}\forall\exists} = \sim_{\text{PTe-tbt,dis}} \subseteq \sim_{\text{PTe-tbt}} \subseteq \sim_{\text{PTe-tbt,} \sqcup \sqcap}$.
- 3. $\sim_{\text{PTe-tbt.dis}} \subseteq \sim_{\text{PRTr.dis}}$.
- 4. $\sim_{PF} \subset \sim_{PTe-tht} \subset \sim_{PTr}$.
- 5. $\sim_{\text{PF}, \sqcup \sqcap} \subseteq \sim_{\text{PTe-tbt}, \sqcup \sqcap} \subseteq \sim_{\text{PTr}, \sqcup \sqcap}$.

All the inclusions above are strict, as shown in Figs. 1(i) and (ii) and Figs. 2(i) to (iii). It is worth noting the isolation of $\sim_{\text{PTe-}\sqcup\square}$, which is incomparable with $\sim_{\text{PRTr,dis}}$, $\sim_{\text{PFTr,dis}}$, $\sim_{\text{PR,dis}}$, $\sim_{\text{PF,dis}}$, $\sim_{\text{PCTr,dis}}$, $\sim_{\text{PRTr,dis}}$, $\sim_{\text{PR,dis}}$, $\sim_{\text{PR,dis}}$, $\sim_{\text{PCTr,dis}}$, $\sim_{\text{PCTr,dis}}$, $\sim_{\text{PR,dis}}$, $\sim_{\text{$

5 Bisimulation Equivalences for NPLTS Models

Bisimulation equivalences capture the ability of two processes of mimicking each other's behavior stepwise. Similar to the trace-based case, given two states there are three different approaches to the definition of these bisimilarities, each following the style of [21] based on equivalence relations. The first approach is to match transitions on the basis of *class distributions*, which means that for each transition of one of the two states there must exist an equally labeled transition of the other state such that, *for every equivalence class*, the two transitions have the same probability of reaching a state in that class. In other words, matching transitions of the two states are related by the fully probabilistic version of bisimilarity (fully matching transitions). The second approach is to consider *a single equivalence class at a time*, i.e., to anticipate the quantification over classes. In this way, a transition departing from one of the two states is allowed to be matched, with respect to the probabilities of reaching different classes, by several different transitions departing from the other state (partially matching transitions). The third approach is to compare only the *extremal probabilities* of reaching each class over all possible transitions labeled with a certain action (max-min-matching transition sets).

Unlike [21], we will consider *groups of equivalence classes* rather than individual equivalence classes. This does not change the discriminating power in the case of the first approach, while it increases the discriminating power thereby resulting in desirable logical characterizations in the case of the other two approaches [8, 32, 31, 5]. Given an NPLTS (S,A,\longrightarrow) and a distribution $\mathscr{D} \in Distr(S)$, in the following we let $\mathscr{D}(S') = \sum_{s \in S'} \mathscr{D}(s)$ for $S' \subseteq S$. Moreover, given an equivalence relation \mathscr{B} over S and a group of equivalence classes $\mathscr{G} \in 2^{S/\mathscr{B}}$, we also let $\bigcup \mathscr{G} = \bigcup_{C \in \mathscr{G}} C$.

Definition 5.1 (*Probabilistic group-distribution bisimilarity* $-\sim_{PB.dis}$ - [28])

 $s_1 \sim_{\text{PB,dis}} s_2$ iff (s_1, s_2) belongs to the largest probabilistic group-distribution bisimulation. An equivalence relation \mathscr{B} over S is a *probabilistic group-distribution bisimulation* iff, whenever $(s_1, s_2) \in \mathscr{B}$, then for each $s_1 \stackrel{a}{\longrightarrow} \mathscr{D}_1$ there exists $s_2 \stackrel{a}{\longrightarrow} \mathscr{D}_2$ such that for all $\mathscr{G} \in 2^{S/\mathscr{B}}$ it holds that $\mathscr{D}_1(\bigcup \mathscr{G}) = \mathscr{D}_2(\bigcup \mathscr{G})$.

Definition 5.2 (*Probabilistic bisimilarity* $-\sim_{PB} - [5]$)

 $s_1 \sim_{PB} s_2$ iff (s_1, s_2) belongs to the largest probabilistic bisimulation. An equivalence relation \mathscr{B} over S is a *probabilistic bisimulation* iff, whenever $(s_1, s_2) \in \mathscr{B}$, then for all $\mathscr{G} \in 2^{S/\mathscr{B}}$ it holds that for each $s_1 \stackrel{a}{\longrightarrow} \mathscr{D}_1$ there exists $s_2 \stackrel{a}{\longrightarrow} \mathscr{D}_2$ such that $\mathscr{D}_1(\bigcup \mathscr{G}) = \mathscr{D}_2(\bigcup \mathscr{G})$.

Definition 5.3 (*Probabilistic* $\Box \Box$ -*bisimilarity* $- \sim_{PB,\Box\Box} - [5]$)

 $s_1 \sim_{\text{PB}, \sqcup \sqcap} s_2$ iff (s_1, s_2) belongs to the largest probabilistic $\sqcup \sqcap$ -bisimulation. An equivalence relation \mathscr{B} over S is a *probabilistic* $\sqcup \sqcap$ -bisimulation iff, whenever $(s_1, s_2) \in \mathscr{B}$, then for all $\mathscr{G} \in 2^{S/\mathscr{B}}$ and $a \in A$ it holds that $s_1 \stackrel{a}{\longrightarrow} \text{iff } s_2 \stackrel{a}{\longrightarrow} \text{ and}$:

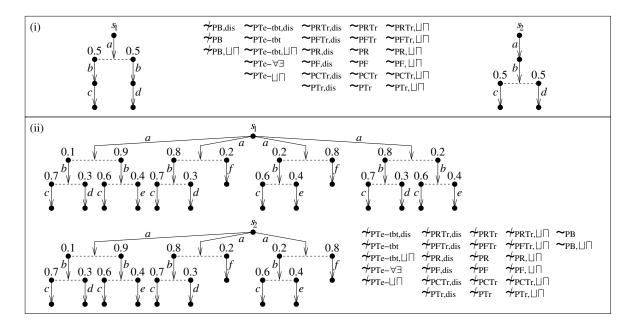


Figure 3: Counterexamples for strict inclusion and incomparability of the bisimulation equivalences

$$\bigsqcup_{\substack{s_1 \xrightarrow{a} \mathscr{D}_1 \\ s_1 \xrightarrow{a} \mathscr{D}_1}} \mathscr{D}_1(\bigcup \mathscr{G}) = \bigsqcup_{\substack{s_2 \xrightarrow{a} \mathscr{D}_2 \\ s_1 \xrightarrow{a} \mathscr{D}_2}} \mathscr{D}_2(\bigcup \mathscr{G})$$

The three bisimulation equivalences defined above are all backward compatible with the bisimulation equivalences respectively defined in [16] for fully nondeterministic processes – which we denote by $\sim_{B,fnd}$ – and in [13] for fully probabilistic processes – which we denote by $\sim_{B,fpr}$.

Theorem 5.4 It holds that:

- 1. $\sim_{PB,dis} = \sim_{PB} = \sim_{PB, \sqcup \sqcap} = \sim_{B,find}$ over fully nondeterministic NPLTS models.
- 2. $\sim_{PB,dis} = \sim_{PB} = \sim_{PB, \sqcup \sqcap} = \sim_{B,fpr}$ over fully probabilistic NPLTS models.

We now investigate the relationships of the three bisimulation equivalences among themselves (first property below) and with the five testing equivalences and the eighteen trace-based equivalences (second property below).

Theorem 5.5 It holds that:

1.
$$\sim_{PB,dis} \subseteq \sim_{PB} \subseteq \sim_{PB,\sqcup\sqcap}$$
.

2.
$$\sim_{\text{PB,dis}} \subseteq \sim_{\text{PTe-tbt,dis}}$$
.

All the inclusions above are strict, as shown in Figs. 1(i) and (ii) and Fig. 3(i). It is worth noting the isolation of \sim_{PB} and $\sim_{PB, \sqcup \sqcap}$, which are incomparable with all the five testing equivalences and all the eighteen trace-based equivalences, as shown in Figs. 3(i) and (ii).

6 Conclusion

We have studied the relationships among the equivalences that stem from three significantly different approaches to the definition of behavioral relations for NPLTS models. The specificity of the three approaches is determined by the way they deal with the probabilities associated with the resolutions of nondeterminism. For each approach, we have considered the families of strong trace-based, testing, and bisimulation equivalences under deterministic schedulers. The relationships among the equivalences for finitely-branching NPLTS models are summarized in Fig. 4. In the spectrum, the absence of (chains of) arrows represents incomparability, adjacency of boxes within the same fragment and double arrows connecting boxes of different fragments indicate coincidence, and single arrows stand for the strictly-more-discriminating-than relation.

Continuous hexagonal boxes contain equivalences studied in the last twenty years [28, 26, 27], which compare probability distributions of all equivalence-specific events. In contrast, continuous rounded boxes contain equivalences assigning a weaker role to schedulers that have been recently introduced in [8, 32, 31, 3, 5], which compare separately the probabilities of individual equivalence-specific events. Continuous rectangular boxes instead contain old equivalences [35, 19, 27, 11] and new equivalences [5] based on extremal probabilities. The only hybrid box is the one containing $\sim_{\text{PTe-V}}$, as this equivalence does not follow any of the three definitional approaches. Finally, dashed boxes contain equivalences defined for the first time in this paper to better assess the different impact of the approaches themselves.

Figure 4 evidences that the top fragment of the spectrum collapses several equivalences, whilst the middle fragment and the bottom fragment do not. Indeed, like in the spectrum for fully probabilistic processes [20, 17], we have that the top variants of ready-trace and failure-trace equivalences and of readiness and failure equivalences respectively induce the same identifications. In contrast, the more liberal variants in the middle fragment and the bottom fragment, which guarantee a higher degree of flexibility in determining the matching resolutions and are in general coarser, do not flatten the specificity of the intuition behind the original definition of the behavioral equivalences for LTS models. Therefore, those two fragments preserve much of the original spectrum of [14] for fully nondeterministic processes. We finally stress again the isolation of \sim_{PB} , $\sim_{PB, \sqcup\sqcap}$, \sim_{PRTr} , \sim_{PR} , $\sim_{PRTr, \sqcup\sqcap}$, and $\sim_{PR, \sqcup\sqcap}$.

As future work, we intend first of all to enrich the spectrum with simulation equivalences. Secondly, we plan to address how the spectrum changes if randomized schedulers are used. Thirdly, we would like to investigate the spectrum of weak behavioral equivalences, for which the choice of randomized schedulers is more appropriate. Finally, it would be interesting to compare the discriminating power of the various equivalences after defining them more abstractly on a parametric model. A suitable framework might be that of ULTRAS [4], as it has been shown to encompass trace, testing, and bisimulation equivalences for models such as labeled transition systems, discrete-/continuous-time Markov chains, and discrete-/continuous-time Markov decision processes without/with internal nondeterminism.

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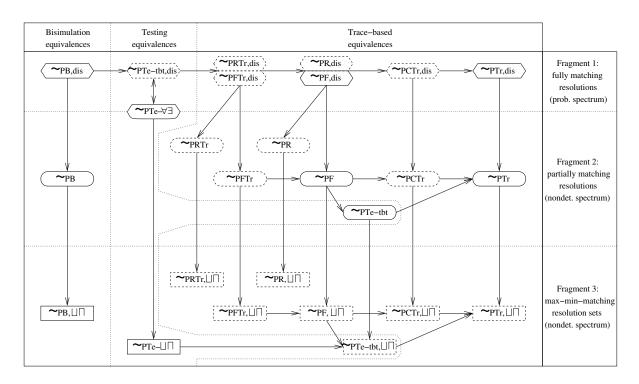


Figure 4: Spectrum of strong behavioral equivalences for NPLTS models (deterministic schedulers)

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